A Reconsideration of the Gibbs' Sampler for Small Area Estimation Models

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The views expressed here are those of the author and not those of the U.S. Census Bureau.

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- Many papers applying it to various statistical models.
- Natural application to linear mixed models, which are popular in small area estimation (SAE).
- Leading MCMC algorithm; produces dependent simulation sequences.

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We shall examine the performance of the Gibbs' sampler versus alternatives for two examples done using JAGS.

Example 1: Fay-Herriot (1979) model

$$y_i|\theta_i \sim N(\theta_i, v_i)$$
 $i=1,\ldots, m=51$

$$\theta_i | \beta, \sigma^2 \sim N(x_i' \beta, \sigma^2)$$

 $p(\beta, \sigma^2) \propto 1$

- y_i = direct survey estimate of population characteristic θ_i for area i
- v_i = sampling variance of y_i (v_i assumed known)
- x_i = vector of regression variables (3 + intercept) for area *i*
- $\beta =$ vector of regression parameters

Example 1: Gibbs' sampler for the FH model

Full conditional distributions for the FH model with $p(\beta, \sigma^2) \propto 1$, $\mathbf{y} = (y_1, \dots, y_m)'$, $\boldsymbol{\theta} = (\theta_1, \dots, \theta_m)'$:

•
$$[\beta|\mathbf{y}, \boldsymbol{\theta}, \sigma^2] = [\beta|\boldsymbol{\theta}, \sigma^2] \sim N(\hat{\beta}, V(\hat{\beta}))$$

where $\hat{\beta} = (X'X)^{-1}X'\boldsymbol{\theta}, V(\hat{\beta}) = \sigma^2(X'X)^{-1} = O(1/m).$

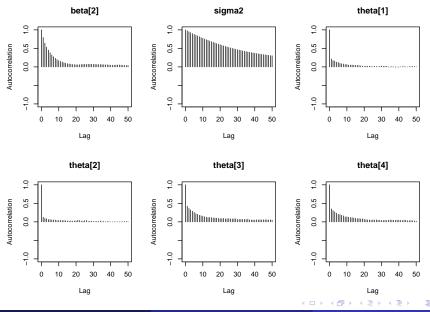
•
$$[\sigma^2 | \mathbf{y}, \boldsymbol{\theta}, \boldsymbol{\beta}] = [\sigma^2 | \boldsymbol{\theta}, \boldsymbol{\beta}] \sim \hat{\sigma}^2 \left(\frac{\chi^2_{m+2}}{m}\right)^{-1}$$

where $\hat{\sigma}^2 = \frac{1}{m} \sum_{i=1}^m (\theta_i - x'_i \boldsymbol{\beta})^2$.

•
$$[\theta_i | y_i, \beta, \sigma^2] \sim N(\hat{\theta}_i, V(\hat{\theta}_i))$$

where $\hat{\theta}_i = h_i y_i + (1 - h_i) x'_i \beta, V(\hat{\theta}_i) = h_i v_i, h_i = \frac{\sigma^2}{\sigma^2 + v_i}$

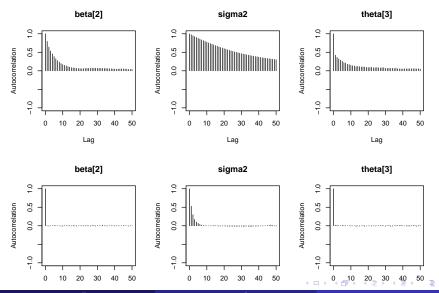
Example 1: FH model, Gibbs' sampler ACFs



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Example 1: FH model, ACFs for Gibbs' sampler and for non-hierarchical model specified to JAGS



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Example 1: Fay-Herriot model

MCMC variance ratio: MC variance of posterior mean for hierarchical specification MC variance of posterior mean for non-hierarchical specification

parameter	β_2	σ^2	$ heta_1$	θ_2	θ_3	
variance ratio	10.7	22.4	4.3	3.6	11.6	

Example 1: Alternative specifications of the FH model

Hierarchical model specification:

$$y_i | \theta_i \sim N(\theta_i, v_i)$$

$$\theta_i | \beta, \sigma^2 \sim N(x'_i \beta, \sigma^2)$$

Non-hierarchical model specification:

$$y_i = \theta_i + e_i \sim N(x_i'\beta, \sigma^2 + v_i)$$

Here we "integrate out the random effects," which avoids the Gibbs' sampler.

Example 2: Bivariate functional measurement error model

For
$$i=1,\ldots,m=51$$
, let $y_i=(y_{1i},y_{2i})'$ and $heta_i=(heta_{1i}, heta_{2i})'$. Then

$$y_i | heta_i \sim N\left(heta_i, \mathbf{V}_i
ight)$$
, $\mathbf{V}_i = \begin{bmatrix} v_{i11} & 0 \\ 0 & v_{i22} \end{bmatrix}$

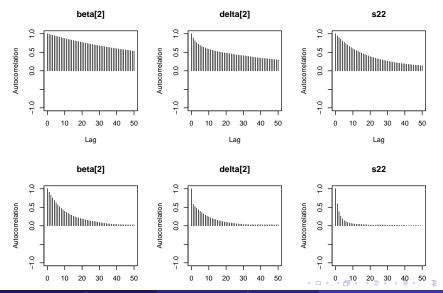
$$\begin{split} \theta_{i1} | x_i, \beta_1, \delta_1 &\sim N \left(\beta_1 x_i + z'_{i1} \delta_1, \sigma_{11} \right) \\ \theta_{i2} | x_i, \beta_2, \delta_2 &\sim N \left(\beta_2 x_i + z'_{i2} \delta_2, \sigma_{22} \right) \\ \text{cov}(\theta_{i1}, \theta_{i2} | x_i, \beta, \delta) &= \sigma_{12} \qquad \rho = \sigma_{12} / \sqrt{\sigma_{11} \sigma_{22}} \end{split}$$

 $X_i | x_i \sim N(x_i, C_i)$ C_i assumed known

Note that

$$E(heta_{ij}|X_i,eta,\delta)=eta_jX_i+z_{ij}'\delta_j$$
 $\operatorname{var}(heta_{ij}|X_i,eta,\delta)=eta_j^2C_i+\sigma_{jj}.$

Example 2: Bivariate functional measurement error model ACFs - Gibbs' sampler vs non-hierarchical spec to JAGS

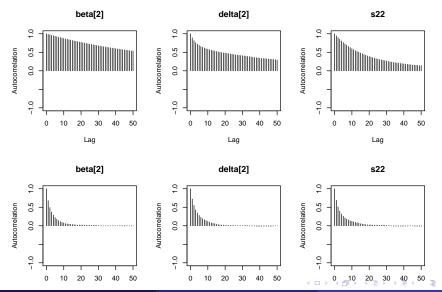


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Example 2: Bivariate functional measurement error model ACFs - Gibbs' sampler vs independence chain



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MCMC variance ratio h/nh: $\frac{MC \text{ variance of posterior mean for hierarchical spec}}{MC \text{ variance of posterior mean for non-hierarchical spec}}$

MCMC variance ratio h/Ind: $\frac{MC}{MC}$ variance of posterior mean for hierarchical spec

parameter	β_2	δ_2	σ_{22}	ρ
variance ratio h/nh				9.8
variance ratio h/Ind	18.2	9.3	4.1	

Why the poor performance of the Gibbs' sampler?

Full conditional distributions for the FH model with $p(\beta, \sigma^2) \propto 1$, $\mathbf{y} = (y_1, \dots, y_m)'$, $\boldsymbol{\theta} = (\theta_1, \dots, \theta_m)'$:

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where $\hat{\beta} = (X'X)^{-1}X'\boldsymbol{\theta}, V(\hat{\beta}) = \sigma^2(X'X)^{-1} = O(1/m).$

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where $\hat{\sigma}^2 = \frac{1}{m} \sum_{i=1}^m (\theta_i - x'_i \boldsymbol{\beta})^2.$

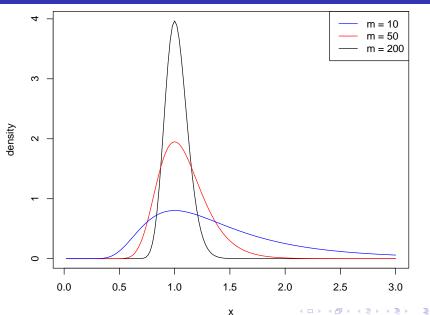
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$$[\theta_i|y_i, \beta, \sigma^2] \sim N(\hat{\theta}_i, V(\hat{\theta}_i))$$

where $\hat{\theta}_i = h_i y_i + (1 - h_i) x'_i \beta, V(\hat{\theta}_i) = h_i v_i, h_i = \frac{\sigma^2}{\sigma^2 + v_i}$.

"Thinning" of the MCMC chains does not solve the dependence problem.

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Densities of m/χ^2_{m+2}



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- Though convenient, the numerical performance of the Gibbs' sampler for mixed linear models, such as models used in small area estimation, is quite poor.
- More efficient algorithms are readily available, including by simply programming mixed linear models in non-hierarchical form in BUGS or JAGS.